UNITED STATES DEPARTMENT OF THE INTERIOR, Oscar L. Chapman, Secretary FISH AND WILDLIFE SERVICE, Albert M. Day, Director

COMPARISON OF YELLOWFIN TUNA OF HAWAIIAN WATERS AND OF THE AMERICAN WEST COAST

BY MILNER B. SCHAEFER



FISHERY BULLETIN 72 From Fishery Bulletin of the Fish and Wildlife Service VOLUME 52

UNITED STATES GOVERNMENT PRINTING OFFICE

WASHINGTON : 1952

For sale by the Superintendent of Documents, U. S. Government Printing Office, Washington 25, D. C.

Price 15 cents

CONTENTS

Introduction	Page 353
On the selection of regression equations	354
Relative growth of Hawaiian yellowfin tuna	358
Comparison of tuna from Hawaii and from the American west coast	359
Fin lengths	359
Head lengths and distances from snout to fin insertions	362
Other dimensions	371
Counts of gill rakers	371
Discussion	371
Literature cited	373

COMPARISON OF YELLOWFIN TUNA OF HAWAIIAN WATERS AND OF THE AMERICAN WEST COAST

By MILNER B. SCHAEFER, Fishery Research Biologist

The yellowfin tuna of the vicinity of the Hawaiian Islands, like the form from the adjacent waters of the American west coast (Schaefer 1948), is here referred to Neothunnus macropterus (Temminck and Schlegel) 1842. As has been pointed out previously (Schaefer and Walford 1950), it is possible that the various Pacific forms, the form from the Indian Ocean, and perhaps also those from the Atlantic, should be considered a single species of world-wide distribution. The data presented herein support such a conclusion. This cannot be finally settled until populations from more places have been carefully studied, particularly a series from the Indian Ocean from which was described the specimen of N. argentivittatus (Cuvier and Valenciennes) 1831, which should be considered the type of this species.

It is also my opinion that the species now referred to the genera Thunnus, Neothunnus, Parathunnus, and Kishinoella should all be referred, as has been done by Fraser-Brunner (1950), to a single genus, Thunnus. However, since this paper is written to compare the yellowfin tuna from the vicinity of the Hawaiian Islands with the form from the waters adjacent to the American west coast in order to settle the question whether they are racially distinct, questions of taxonomy, synonymy, and nomenclature will be passed over at this time, and for convenience both forms will be referred to the commonly accepted name N. macropterus.

The yellowfin tuna is the object of an extensive and intensive fishery along the American west coast from California to the Galapagos Islands. In the Hawaiian Islands there exists a minor fishery that promises to be expanded in the near future to encompass other islands of the mid-Pacific and to increase in intensity in the presently exploited region. Whether the population of the Hawaiian region is part of the same stock of fish as that fished along the American west coast, or is an independent stock, is a question of con-

. . .

siderable practical importance: if they are the same stock, the new fisheries would merely add to the strain on the stock already being exploited; if they are independent, there is being tapped an essentially virgin resource.

Schaefer (1948) has published measurement data and counts of denumerable characters on yellowfin tuna from the waters of the Pacific near Costa Rica. Godsil (1948) has published the measurements of a few selected dimensions taken from a very large number of specimens from several sampling localities, extending from the tip of lower California to Panama. Godsil and Byers (1944) have also published gill-raker counts of value to the present study. Those data and those presented herein from the Hawaiian Islands are directly comparable, having been taken in the same manner. Details of measurement methods are given in the papers cited and by Marr and Schaefer (1949). Measurements were made by several field assistants, but all followed identical procedures.

For this study, Hawaiian yellowfin tuna were measured during 1949, between February 21 and September 28. They were selected to give as even a representation as practicable of all sizes of fish available. All specimens were fresh and recently landed from commercial fishing vessels. Most specimens were measured at the Honolulu fresh-fish wholesale auction market, not only a very convenient place to work but almost ideal from a sampling standpoint.

The fish handled there are caught by flaglines which, by the nature of their operation, sample the fish population very widely. Description of the fishery and the method of handling and marketing the fish will be found in June (1950). Smaller sizes of yellowfin tuna, under about 80 cm. in total length, are seldom taken by the flag-line fishery. These small fish are frequently taken by pole-and-line fishing, in the same manner as on the American west coast,

.

Rejected

incidental to fishing for skipjack (Katsuwonus pelamis). Specimens of the small sizes were mostly obtained, therefore, from landings at the local tuna cannery, where most of the skipjack catch is landed, particularly during the summer season of good catches. These fish are landed fresh soon after being caught, and are thus comparable to the specimens from the flag-line fishery. The original data on the 203 Hawaiian yellowfin tuna employed in this study are tabulated in table 1. All length measurements are in millimeters, taken as described by Marr and Schaefer (1949). Weights were taken in pounds, because at the auction market the fish were weighed by commercial scales graduated in pounds. Blanks in the table indicate that the measurement or count was not taken on the particular specimen. In addition, a few of the tabulated values were omitted from the analyses, because they were found to deviate more than three standard deviations from the appropriate regression line and seemed probably to be recording errors. These values were as follows:

	value
1670-mm. specimen, snout to insertion first dorsal_	423
1780-mm. specimen, snout to insertion first dorsal	446
1780-mm. specimen, snout to insertion second	
dorsal	835
1464-mm. specimen, snout to insertion anal	767
1629-mm. specimen, body depth	454
1333-mm. specimen, longest dorsal finlet	34
1259-mm. specimen, length first dorsal spine	97
1397-mm. specimen, length first dorsal spine	129
969-mm. specimen, diameter of iris	26
1605-mm. specimen, diameter of iris	52

Many of the routine computations involved in the analysis of the Hawaiian data, reanalysis of American-west-coast data, and comparison of the two, were performed by Dorothy Dung, whose assistance is gratefully acknowledged.

ON THE SELECTION OF REGRESSION EQUATIONS

It is characteristic of many animals—perhaps of all—that the various parts of the body grow at different rates, so that as the organism increases in size the ratio of one dimension to another changes. For yellowfin tuna this has been demonstrated by Godsil (1948), Schaefer (1948), and Schaefer and Walford (1950). Since this is the case, one cannot use the measurement ratios normally employed in systematic ichthyology for comparing samples of tunas from different places; except in the trivial case where the fish from the two places are of exactly the same size, because differences connected with size could be confused with differences in form of fish of the same size.

In order to avoid this difficulty, the authors of the papers cited above have based their comparisons of samples on the comparison of the regression of one dimension on that of another (usually total length), taken as a measurement of over-all size. This procedure is also employed in the present paper. It may be noted that the efficiency of sampling may be much improved over simple random sampling in such circumstances by selecting the specimens according to total length (the independent variate) to give an even representation of all sizes available so far as is practical; such a sampling scheme was employed in obtaining the data for table 1.

The comparison of body form among fish populations by comparison of regressions would be a simple and straightforward process if the relations between the body dimensions corresponded exactly to the straight lines or simple curves that must be employed in such analyses. Unfortunately, they do not and this may lead to some confusion in the analysis, particularly in situations where one is dealing with small differences and large numbers of specimens. Over restricted ranges of sizes at least, the dimensions of some body parts relative to others seem to be sufficiently well approximated by straight lines (Schaefer 1948, Schaefer and Walford 1950). Large samples of the same size range of the same populations may reveal, however, that regression curves of slight curvilinearity give a better fit to the data, as Godsil (1948) has found for certain dimensions of the American-west-coast yellowfin.

In other cases, such as the fin lengths of yellowfin tuna, the regressions are very strongly curvilinear but may, in some cases at least, be transformed by the allometry equation or other transformation to a linear or nearly linear relation, as has been done in my papers above cited. Whatever the equation employed, however, it is necessary to bear in mind two things. First, the relation employed in the analysis (the mathematical model of the true relation between variables), be it linear or otherwise, is only an approximation to the true relation and as such does not completely eliminate the effect of size of organism on the character being compared. Second, there sometimes occur rather marked changes in growth rate of one part relative to another at certain sizes, so that a regression which over a considerable range may be represented by a particular equation may not be so represented at all when the range is slightly extended. Indeed, as has been shown by Martin (1949), there seem to be sharp inflection points in the relativegrowth curves of several fish species. The avoidance of misleading conclusions demands that these matters be kept in mind in analyses of morphometric data.

TABLE	1.—Morphometric	measurem e nts	and	counts	for	Yellowfin	tuna	(Neothunnus	macropterus)	from	the	Hawaiian
•				Islands	Fel	b. 21–Sept.	28, 1	949		-	•	

Total length	Weight	Head length	Snout to insertion first dorsal	Snout to insertion second dorsal	Snout to insertion anal	Snout to insertion ventral	Greatest depth	Taken at dorsal spine No	Length pectoral fin	Length first dorsal (first spine)	Length second dor- sal	Length anal	Length longest dor- sal finlet	Longest dorsal, No	Diameter of iris	Length of maxil- lary	Number of spines first dorsal	Number of dorsal finlets	Number of anal finlets	Number of gill rakers	Rex
451 mm	$\begin{array}{c} L_{4,0} \\ 5_{5,0} \\ 5_{5,0} \\ 5_{5,5} \\$	$\begin{array}{r} \mathcal{M}m.\\ 128\\ 135\\ 137\\ 132\\ 132\\ 132\\ 132\\ 132\\ 132\\ 132\\ 133\\ 137\\ 128\\ 137\\ 138\\ 137\\ 140\\ 141\\ 142\\ 144\\ 144\\ 144\\ 144\\ 144\\ 144$	Mm. 143 143 152 152 152 153 148 155 155 155 155 155 155 155 155 155 15	$\begin{array}{c} \lambda m. \\ 249 \\ 255 \\ 262 \\ 263 \\ 264 \\ 265 \\ 264 \\ 265 \\ 264 \\ 274 \\ 266 \\ 265 \\ 264 \\ 277 \\ 275 \\ 277 \\ 281 \\ 283 \\ 282 \\ 283 \\ 282 \\ 283 $	Mm. 272 281 282 284 287 292 294 287 295 293 296 303 302 296 303 303 302 296 303 303 302 296 303 303 302 297 303 308 303 309 303 301 302 303 303 303 303 303 303 3115 315 325 333 310 325 325 556 550 550 550 550 557 567 568 543 500 555 553 567 568 574 567 568 563 605 6036	Mm. 145 146 133 133 148 153 154 154 155 157 156 157 157 166 159 171 177 164 168 163 163 163 163 163 163 163 163 163 163	$\begin{array}{c} Mm. \\ 1144\\ 125\\ 138\\ 125\\ 138\\ 123\\ 123\\ 123\\ 123\\ 123\\ 123\\ 123\\ 123$	77677767796898767667076710898989896878879677787 77884877787977977858789999	$\begin{array}{c} Afm. \\ 125\\ 134\\ 132\\ 133\\ 143\\ 143\\ 143\\ 143\\ 144\\ 143\\ 144\\ 144$	Mm. 63 63 54 54 56 57 58 58 57 58 58 57 58 58 59 62 62 58 58 59 62 62 62 58 58 59 60 63 58 58 59 60 63 74 73 88 88 88 80 60 63 74 73 83 83 80 67 74 73 83 83 80 67 74 73 83 83 80 67 74 73 83 83 80 60 60 60 60 60 60 60 60 60 6	Mm. Mm. 50 48 50 51 55 55 55 55 55 56 57 58 59 66 65 66 68 68 68 68 68 109 137 166 1839 190 202 203 203 203 204 324 324	Mm. 44 388 494 479 494 479 495 535 600 586 605 566 605 565 612 625 555 600 555 600 555 600 655 660 655 660 655 660 655 660 655 660 625 535 660 655 660 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 535 600 625 727 7177 7177 7177 7175 2211 2200 5215 2205 2215 2205 2215 2205 2215 2205 2215	$\mathcal{M}_{m.}$ 14 17 16 16 17 16 16 17 17 16 17 17 16 17 17 16 17 17 16 17 17 16 17 17 16 17 17 16 17 17 16 19 19 19 19 19 19 19 19 19 19 19 19 19	50566055566555666655566665566555 5666555 566555 5666666	n.325425533855234488552528655752686758275526882738273883333343286 348643335 353334333323555445888555333453885553334538553334538553334538555333453555333453555533345555333455553334555553334555553334555553334555553334555553334555553334555553334555555	Mn, 50 50 50 50 50 50 50 50 50 50 50 50 50 5	$\begin{array}{c} 14\\ 13\\ 14\\ 14\\ 14\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13$	9+1-2-2-2-2-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1	\$\$\$1,227,537,557,577,777,777,777,777,777,777,77	$\begin{array}{c} 10+20\\ 10+21\\ 10+21\\ 9+22\\ 9+$	MFFMMFM - (?) MMF MMM(?) MMM(?) MMM M MMF F

FISHERY BULLETIN OF THE FISH AND WILDLIFE SERVICE

																		_	_		
Total length	Weight	Hoad length	Snout to insertion first dorsal	Snout to insertion second dorsal	Snout to insertion anal	Snout to insertion ventral	Greatest depth	Taken at dorsal spine No. —	Length pectoral fin	Length first dorsal (first spine)	Length second dor- sal	Length anal	Leugth longest dor- sal finlet	Longest dorsal, No. –	Diameter of iris	Length of maxil- lary	Number of spines first dorsal	Number of dorsal finlets	Number of anal finlets	Number of gill rakers	Sex
1,206 mm 1,217 mm 1,238 mm 1,239 mm 1,239 mm 1,240 mm 1,255 mm 1,256 mm 1,256 mm 1,258 mm 1,258 mm 1,259 mm 1,278 mm 1,278 mm 1,289 mm 1,289 mm 1,289 mm 1,289 mm 1,281 mm 1,282 mm 1,283 mm 1,323 mm 1,331 mm 1,332 mm 1,333 mm 1,331 mm 1,332 mm 1,333 mm 1,339 mm 1,330 mm 1,331 mm 1,332 mm 1,333 mm 1,339 mm 1,330 mm 1,330 mm 1,330 mm 1,340 mm		$\begin{array}{c} \mathcal{M}m, \\ 3090\\ 313\\ 310\\ 313\\ 310\\ 313\\ 311\\ 307\\ 311\\ 311\\ 307\\ 311\\ 311\\ 311\\ 311\\ 311\\ 311\\ 311\\ 31$	$\begin{array}{l} \mathcal{M}m, & 342\\ \mathcal{M}m, & 346\\ 348\\ 348\\ 348\\ 348\\ 343\\ 351\\ 356\\ 356\\ 356\\ 356\\ 356\\ 356\\ 356\\ 356$	$\begin{array}{c} \mathcal{M}m,\\ 605\\ 621\\ 621\\ 621\\ 621\\ 622\\ 622\\ 632\\ 632\\ 645\\ 652\\ 9632\\ 645\\ 6529\\ 632\\ 645\\ 6529\\ 632\\ 645\\ 643\\ 648\\ 645\\ 643\\ 648\\ 645\\ 668\\ 666\\ 666\\ 666\\ 666\\ 666\\ 666\\ 66$	Mm. 685 683 694 6843 694 6893 694 700 720 720 720 7216 729 723 746 7407 757 746 740 757 7733 746 740 757 7732 7746 752 7757 752 7747 752 7759 759 7762 777 7803 786 7804 7804 7805 786 7806 7999 7807 792 7777 780 7808 7806 7809 8307 7808 850 8300 8323 8311 764 8057 836 8378 837 8380 852 8378 852	$\begin{array}{c} Mm.\\ 346\\ 851\\ 358\\ 358\\ 358\\ 355\\ 357\\ 370\\ 355\\ 356\\ 365\\ 375\\ 356\\ 365\\ 372\\ 238\\ 366\\ 363\\ 377\\ 371\\ 371\\ 371\\ 371\\ 381\\ 381\\ 373\\ 381\\ 381\\ 375\\ 376\\ 376\\ 376\\ 376\\ 376\\ 376\\ 376\\ 376$	Mm. 2299 3000 3000 3000 3000 3000 3000 3000	$\begin{array}{c} 979910\\ 9799877910\\ 10779988611\\ 777899697711910\\ 9777910968778977791096877897779109688881098881098881091117999888867111019977791098888109111977791091111111111111111$	$\begin{array}{c} Mnn.\\ 3360\\ 3360\\ 3360\\ 3421\\ 3441\\ 3441\\ 3441\\ 3411\\ 341\\ 3451\\ 3450\\ 3450\\ 349\\ 349\\ 349\\ 349\\ 349\\ 349\\ 349\\ 349$	$\begin{array}{c} Mm. \\ 133 \\ 146 \\ 134 \\ 131 \\ 131 \\ 131 \\ 131 \\ 131 \\ 131 \\ 131 \\ 131 \\ 131 \\ 142 \\ 149 \\ 977 \\ 143 \\ 142 \\ 149 \\ 977 \\ 143 \\ 142 \\ 149 \\ 197 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 134 \\ 141 \\ 132 \\ 137 \\ 132 \\ 137 \\ 162 \\ 157 \\ 139 \\ 146 \\ 155 \\ 157 \\ 139 \\ 146 \\ 155 \\ 157 \\ 139 \\ 147 \\ 138 \\ 129 \\ 107 \\ 151 \\ 156 \\ 157 \\ 157 \\ 139 \\ 161 \\ 155 \\ 161 \\ 161 \\ 161 \\ 161 \\ 161 \\ 161 \\$	$\begin{array}{c} Mm.\\ 278\\ 278\\ 278\\ 345\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 278\\ 276\\ 273\\ 270\\ 273\\ 273\\ 273\\ 273\\ 273\\ 273\\ 273\\ 273$	$\begin{array}{c} \mathcal{M}m.\\ 315\\ 317\\ 327\\ 327\\ 3327\\ 3327\\ 3327\\ 3327\\ 3327\\ 3327\\ 3322\\ 3333\\ 3333\\ 33333\\ 333333\\ 3333333\\ 33333333\\ 333333333333333333333333333333333333$	$\begin{array}{c} \mathbf{Mm.} \\ 433 \\ 443 \\ 445 \\ 544 \\ 445 \\ 544 \\ 455 \\ 544 \\ 455 \\ 544 \\ 455 \\ 554 \\ 556 \\ 555$	65667665575565050408556808566074555664878656086555665646608665556608665556565646608665605745556545647658766	$\begin{array}{c} Mm.88737773333839938399383939143838377373338399383993839391438383773733383839143838373737333839143838373733383914383838373733383939143838383737333839391438383837373338393914383838373733383939143838383737333839391438383837373338393914383838373733383937333839373338383737333838373733383837373338383737333838373733383833373733383837373338383837373338383337373338383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733383833373733333333333333333333$	$\begin{array}{c} Mm. \\ 1122\\ 1127\\ 1120\\ 1117\\ 120\\ 1127\\ 1217\\ 1217\\ 1217\\ 1217\\ 1217\\ 1217\\ 1217\\ 1218\\ 121\\ 124\\ 121\\ 121\\ 121\\ 121\\ 121\\ 121$	$\begin{array}{c} 13 \\ 14 \\ 14 \\ 14 \\ 13 \\ 13 \\ 13 \\ 13 \\$	<u>1111111111111111111111111111111111111</u>	┰┰┰┰┰┰┰┰┰┰┰┰┰┰┰┰┰┲╕╕╕┲╕╕┙┙┙┙┙┙┙┙┙┙┙┙┙┙┙		F F F F F M F F M F F M F F M F F M F F M F F M F F M M F F M M M M M M M M M M M M M

TABLE 1.—Morphometric measurements and counts for Yellowfin tuna (Neothunnus mac Islands Feb. 21-Sept. 28, 1949—Continued	icropterus) from	the Hawaiian
--	------------------	--------------

COMPARISON OF YELLOWFIN TUNA OF HAWAIIAN WATERS AND THE AMERICAN WEST COAST 357

·																	<u> </u>				
Total length	Weight	Head length	Snout to insertion first dorsal	Snout to insertion second dorsal	Snout to insertion anal	Snout to insertion ventral	Greatest depth	Taken at dorsal spine No. —	Length pectoral fin	Length first dorsal (first spine)	Length second dor- sal	Length anal	Length longest dor- sal finlet	Longest dorsal, No	Diameter of fris	Length of maxil- lary	Number of spines first dorsal	Number of dorsal finlets	Number of anal finlets	Number of gill rakers	Sex
1,611 mm 1,614 mm 1,621 mm 1,620 mm 1,620 mm 1,623 mm 1,635 mm 1,638 mm 1,648 mm 1,648 mm 1,648 mm 1,648 mm 1,668 mm 1,668 mm 1,668 mm 1,668 mm 1,668 mm 1,670 mm 1,670 mm 1,673 mm 1,674 mm 1,673 mm 1,673 mm 1,700 mm 1,710 mm 1,710 mm 1,711 mm	Lb. 163 201 1766 202 201 202 202 202 202 202 202	Mm, 303 3939 405 415 411 412 403 399 401 404 403 404 403 404 403 404 403 404 403 404 403 404 403 404 403 404 403 403	Mm. 432_445 445 445 459 470 447 447 447 447 447 447 447 447 447	Afm. 796 811 822 829 829 810 811 813 813 813 813 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 812 813 813 813 813 813 814 815 815 815 815 815 815 815 815 815 815	Mm. 883 877 902 907 905 901 890 890 901 915 919 919 919 919 919 919 91	M m. 445 440 445 452 457 458 457 458 457 458 457 458 458 458 458 458 458 458 458 458 458	$\begin{array}{c} Mn,\\ 393\\ 440\\ 452\\ 453\\ 452\\ 453\\ 452\\ 453\\ 452\\ 453\\ 452\\ 453\\ 452\\ 453\\ 453\\ 453\\ 453\\ 453\\ 453\\ 453\\ 453$	$\begin{array}{c} 7 & 8 \\ 9 \\ 107 \\ 7 \\ 8 \\ 9 \\ 107 \\ 118 \\ 8 \\ 9 \\ 8 \\ 7 \\ 8 \\ 8 \\ 100 \\ 11 \\ 7 \\ 6 \\ 7 \\ 7 \\ 7 \\ 6 \\ 8 \\ 10 \\ 11 \\ 19 \\ 9 \\ 8 \\ 8 \\ 110 \\ 9 \\ 9 \\ 10 \\ 9 \\ 10 \\ 9 \\ 10 \\ 10$	Afm. 366 400 380 380 380 380 390 412 391 417 440 390 391 417 440 390 391 392 393 412 382 380 394 395 412 382 430 384 381 432 382 431 442 420 411 420 411 420 411 420 411 420 411 410	Mrm. 184 184 184 185 195 203 195 195 195 195 195 195 195 195	M m. 3807 567 764 554 453 454 453 454 453 454 453 454 453 554 651 154 4590 564 4590 564 4590 565 567 760 552 770 552 770 552 777 572 777 577 5	Mm. 664 627 745 582 702 672 745 582 702 745 582 702 745 510 720 725 720 725 720 725 638 4 646 4 510 720 725 637 4 510 6 6 6 6 6 6 6 6	Mm. 62 65 63 63 63 63 63 63 63 64 60 66 67 70 61 61 62 63 63 66 66 67 61 61 62 63 63 66 66 66 67 61 61 62 63 63 62 65 63 63 62 63 63 63 63 63 62 63 63 63 63 63 63 63 63 63 66 66 67 61 61 62 63 63 66 66 67 61 61 62 63 63 66 66 66 67 61 61 62 63 66 66 66 67 61 61 62 63 66 66 66 67 61 61 62 63 66 66 66 66 67 61 62 63 66 66 66 66 66 66 66 66 66 66 66 66	65655576566467656666685588665555656766656556656656	M45 444 45 52 45 54 45 54 54 54 54 54	Mm. 156 153 157 158 157 158 154 154 153 155 154 154 155 154 155 154 155 155 155	14 14 13 13 13 13 13 13 13 13 13 13 13 13 13	<u> </u>		9+2121 9+	M M M M M M M M M M M M M M M M M M M

 TABLE 1.—Morphometric measurements and counts for Yellowfin tuna (Neothunnus macropterus) from the Hawaiian Islands Feb. 21-Sept. 28, 1949—Continued

Godsil (1948), whose work will be discussed subsequently, has found that a curvilinear equation fits the regressions on body length of the distances from the tip of the snout to various fin insertions and head length rather better than a linear one. He also discovered that when he fitted regression equations of the selected type to each of several samples from the same region, and also fitted an equation of this same type to the pooled data of all such samples, the individual regressions differed from the regression for the pooled data to a greater extent than might be expected from purely random variation. This he attributed to a lack of "biological homogeneity" (which he contrasts to "statistical homogeneity") within the stock of fish sampled, arising from incomplete mixing of fish from different spawning grounds. This may indeed be true. A rather simpler explanation is that the small differences he found between regressions among the samples from the same region are due to rather great differences in size composition of the several samples and the necessarily approximate nature of the regression equations employed. Whatever the cause, it is necessary to recognize that such differences can and do arise and to take suitable account of them where required, both in the sampling and in the subsequent analysis. By drawing samples widely from many different schools within the region to be studied, one minimizes for purpose of comparison the effects, if any, of lack of "biological homogeneity" by including in the variance of the sample any differences between subdivisions of the population with different genetic histories. By comparing only samples of the same size range from different regions, one will tend to reduce the apparent difference due to the failure of the regression equation employed to completely correct for differences in size composition of the samples.

There is probably no purely routine method of analysis which may be safely employed in comparing body dimensions of tunas from different regions. The selection of regression equations, and the application of other statistical techniques, should be undertaken with proper consideration of the particular data at hand, the hypotheses regarding it that are to be tested, and the precision required in each particular case.

RELATIVE GROWTH OF HAWAIIAN YELLOWFIN TUNA

Schaefer (1948) and Schaefer and Walford (1950) fitted linear regression lines to head length and distances from tip of snout to insertions of the first dorsal, second dorsal, anal, and ventral fins plotted against total length for yellowfin tuna from the west coast of Central America and from the Atlantic coast of Africa. Godsil (1948) found more extensive data on the same dimensions of yellowfin from the American west coast to be better fitted by a regression line of slight curvilinearity. To the Hawaiian data have been fitted linear regressions, the constants for which are given in table 2, as well as curvilinear regressions of the type selected by Godsil. Equations for the latter and corresponding standard errors of estimate (s) about them-are as follows:

Head length	¥=	69. 54+0.	20805x	15419/3	s= (3. 02
Snout to insertion first dorsal	y =	SO. 34+0.	2286Sz	-16997/3	s= 7	7.77
Snout to insertion second dorsal	y =	17.28+0.	482267-	+11445/ <i>x</i>	8=1(). 94
Snout to insertion ventral	y ==	78.87+0.	23340x	-16778/3	8= 7	7.96
Snout to insertion anal	y =	109. 92+0.	49037.5	-25129/x	s= \$	9. 32

Over the range of sizes in our sample, the curvilinear regressions result in slightly smaller variances about them than the linear regressions; but, as may be seen from the above equations or from the graphs in the next section (figs. 6-10), the differences between these curves and straight lines are slight. Indeed, for snout to second dorsal insertion the slight curvature of the regression is opposite in direction to those fitting the data of other dimensions and to that of Godsil for his American-west-coast fish (fig. 8). Furthermore, the difference between the linear and curvilinear regressions for this dimension is, for the Hawaiian data, such as might arise by chance alone in between 1 in 20 and 1 in 100 cases.

The relations between body depth and total length, diameter of iris and head length, and length of maxillary and head length seem to be well described by linear regressions over the entire size range. The statistics of these regressions are tabulated in table 2.

In each of these cases where linear regressions fit the data, the y intercept of the regression line differs significantly from zero. Furthermore, except for depth of body on total length and length' of maxillary on head length, the difference is sufficiently great that the expression as ratios of the relation between variables would result in a considerable error from this source. This

TABLE 2.—Statistics of linear regressions of measurements of Hawaiian N. macropterus

All logarithms are to base 10. N=number in sample.

 $\bar{x}, \bar{y} = \text{means of } x \text{ and } y.$

 Sx^2 , Sy^2 , Sxy are sums of squares and products of deviations from the means x, y.

 $b = \frac{Sxy}{Sx^2}$ = regression coefficient of y on x.

 $s^{3/2} = \frac{Sy^3 - b^2 Sx^3}{27 - a^2}$ = estimate of variance about regression line.

Independent variable x	Dependent variable y	N	Ŧ	ÿ	Sr ²	Sy2	Szy	b	. 8
Total length Do. Do.	Head length Snout to insertion first dorsal Snout to insertion second dorsal Snout to insertion ventral. Snout to insertion anal. Greatest body dopth. Diameter of iris. Length of maxillary. Longth pectoral. Log length second dorsal ¹ Log length anal ¹ Log length first dorsal spine. Log length longest dorsal finlet Weight in pounds.	203 201 202 203 202 202 203 203 203 203 172 188 198 203	1247 1242 1244 1247 1246 1245 315 314 3.06448 3.13093 3.13093 3.13093 3.07795 3.06657 3.06566	$\begin{array}{r} 314\\ 34b\\ 623\\ 354\\ 697\\ 310\\ 37.6\\ 121.8\\ 324\\ 2.55442\\ 2.59632\\ 2.12768\\ 1.62583\\ 1.82955\end{array}$	32, 985, 274 32, 516, 976 32, 699, 372 32, 985, 374 32, 937, 786 32, 835, 336 1, 667, 677 1, 688, 191 6, 53772 1, 54087 5, 42218 6, 4490 6, 47257	$\begin{array}{c} 1, 688, 363\\ 2, 016, 613\\ 7, 221, 223\\ 2, 118, 502\\ 8, 906, 702\\ 2, 162, 089\\ 9, 005\\ 244, 663\\ 1, 617, 580\\ 8, 11623\\ 8, 66325\\ 5, 30176\\ 7, 67031\\ 58, 24926\\ \end{array}$	7. 443, 781 8, 071, 090 15, 340, 597 8, 331, 748 77, 108, 301 8, 363, 540 119, 469 640, 453 8211, 2003 3. 41003 3. 52758 5. 30154 6. 94380 19. 39100	0. 22567 . 24821 . 46914 . 25259 . 51941 . 25469 . 07164 . 37937 491. 93 2. 21305 2. 28934 . 97775 1. 07664 2. 99587	6.51 8.17 11.03 8.34 10.14 12.64 1.51 2.90 13.73 .0579 .0588 .02530 .03146 .02793

¹ Only specimens 600 mm. and over in total length.

result is similar to that obtained from Central American and African yellowfin tuna (Schaefer 1948, Schaefer and Walford 1950) and illustrates again the generalization that, owing to differential growth rates, comparison of dimensions expressed as ratios is invalid for yellowfin tuna.

Also similar to previous Central American and African results, is the finding that the growth of the pectoral fin of Hawaiian yellowfin tuna is such that over the entire range of sizes available in our sample, the relation between length of fin and total length is well described by the equation

$$y = 491.9 \log x - 1184$$
,

a linear regression giving a good fit to the length of fin plotted against logarithm of total length. The regression statistics are given in table 2.

For Central American and African fish, the lengths of second dorsal and anal fins plotted against total length were found to be fitted by an equation of the type $y=ax^{b}$, so that a linear regression was obtained by plotting logarithms of fin length against logarithms of total length. The sizes of fish involved were from about 50 cm. to 160 cm. in total length for the fish from both regions. For Hawaiian vellowfin tuna, a linear relation between logarithm of fin length and logarithm of fish length provides a fairly good fit over the range of sizes 60 cm. to 178 cm., but when smaller sizes are included, the regression is obviously curvilinear (fig. 2 and 3). Linear-regression equations were fitted, for comparative purposes, only to the data for fish 60 cm. and over in total length, the results being tabulated in table To provide a reasonable fit to the data for 2. all sizes, however, the second-degree polynomials illustrated in the figures were fitted, the equations being, for logarithms of length of second dorsal (y_1) on logarithm of total length (x_1) .

 $y_1 = 7.64965 - 5.59555x_1 + 1.26613x_1^2$ s = .05238

and for logarithm of length of anal (y_1) on logarithm of total length (x_1)

$$y_1 = 4.79192 - 3.82511x_1 + 0.99707x_1^2$$
 $s = .03607$

It is obvious that the relative rates of growth of the second dorsal and the anal fins accelerate very rapidly with increase in size of fish, the large fish having, relatively, enormously longer fins.

The equation $y=ax^b$ was found to provide a good fit to our Hawaiian data over the entire

998351-52-2

range of sizes for length of longest dorsal spine (the first spine in each specimen) and length of longest dorsal finlet relative to total length, the logarithms of the dimensions plotted against logarithm of total length being well fitted by linear regressions, the constants for which are given in table 2. In previous studies of Costa Rican and African fish, linear regressions were found adequate for these relations over the size range 50 cm. to 160 cm., and for only that range of sizes it would be difficult to perceive that the allometry equation provides a better fit to the Hawaiian data. The availability of a longer range of sizes from Hawaiian waters made it possible to observe the slightly curvilinear nature of the relation. How little it differs from a straight line may be seen from the closeness to unity of the values of btabulated in table 2 for these regressions.

The weight of Hawaiian yellowfin varies almost exactly as the cube of the length, the relation between length in millimeters (x) and weight in pounds (y) being expressed by the equation

$\log y = 2.996x - 7.35477$

COMPARISON OF TUNA FROM HAWAII AND FROM THE AMERICAN WEST COAST

Fin lengths

The most outstanding differences revealed by this study between yellowfin tuna from Hawaii and those from waters off Costa Rica are the relative lengths of the pectoral, second dorsal, and anal fins. There seem also to be small but dependable differences in length of longest dorsal spine and length of longest dorsal finlet.

Figure 1 illustrates the relation between length of pectoral fin and total length for Hawaiian and Costa Rican fish. The points plotted in this figure, and in the other figures in this paper, do not represent individual fish but are the mean values of the two variables for each 10-cm. size category. This method of plotting recommends itself because the data for individual fish are too numerous to be clearly depicted. It has also the advantage of making possible a visual comparison of mean values of the dimension under consideration for fish of each single 10-cm. size category from the two populations. The inherent disadvantage is, of course, that each point does not represent the same number of fish, so that their positions are of varying degrees of reliability. The regression



FIGURE 1.—Relations between length of pectoral fin and total length. Open circles and fine line represent Costa Rican data. Solid circles and heavy line represent Hawaiian data.

lines depicted in the figures were in every case fitted to the original data and not to the class means.

As may be seen from figure 1, the pectoral fins of Hawaiian yellowfin tuna, over the size range considered, are on the average longer than those of Costa Rican fish, and the difference increases as the size of fish increases. No elaborate statistical analysis is required to show that these samples cannot be considered as arising from the same population. If inspection of the figure itself is not sufficiently convincing, a very simple test suffices to show that the probability of the two samples arising by random sampling from a single population is very small, regardless of whether or not the growth law on the basis of which the regressions were calculated is exactly correct. Under the hypothesis that the Costa Rican sample was drawn from the same population as the Hawaiian sample, we should expect the points for Costa Rican fish to be half the time above and half the time below the corresponding values predicted from the Hawaiian sample. For each size class, the Costa Rican value falls below the value which would be expected on the basis of the Hawaiian sample. The probability of this occurring by chance alone for all 10 Costa Rican points is $(\frac{1}{2})$ ¹⁰ or 1 chance in 1024; it is, then, most unlikely.

In figure 2 are plotted values of logarithm of length of second dorsal fin against logarithm of total length. This transformation yields a linear regression for the Costa Rican sample, the fish in which are from 54 cm. to 157 cm. in total length. Similarly, the Hawaiian data for fish 62 cm. and over in total length are rather well fitted by a linear regression, as shown in the figure (we have no Hawaiian specimens between 54 cm. and 62 cm.). We have also plotted in the figure the second-degree polynomial that fits the Hawaiian data for all sizes of fish in our sample. It is obvious, whichever regression we employ for the Hawaiian fish, that the second dorsal fins of yellowfin tuna from waters of the Hawaiian Islands grow. relative to total length, faster than those of yellowfin tuna from waters off Costa Rica. The difference in fin lengths is small at smaller sizes of fish, but increases with size of fish until among large fish the difference is very striking.

As may be seen from figure 3, the same situation obtains for the length of anal fin relative to total length. As has been reported for Costa Rican fish and African fish, the variability of fin lengths of second dorsal and anal fins, even on a logarithmic scale, is not entirely independent of size of fish, but tends to be greater at larger sizes. For this reason the values of s for the corresponding equations in table 2 and on page 359 are average values, and will be a little too small at large fish sizes and too large at small sizes.

Comparison of the linear regressions of figures 2 and 3 may be made by means of analysis of covariance (Kendall 1946, p. 237 *et seq.*); or, without reference to regression equations, we may simply compare the mean values of the several size classes and, following the same sort of reasoning as above in the case of the pectoral fin, arrive at the conclusion that the probability of the samples being drawn from a single population is very small. The first dorsal spine was the longest on each of the 188 specimens for which this character was measured. As noted on page 359, a linear regression did not provide a good fit to the original data, compared with a linear regression fitted to the logarithms of the variables. The latter is plotted in figure 4. It was found that the same transformation applied to the Costa Rican data, yielded a linear regression with a slightly improved fit to those data also (Schaefer 1948 fitted a linear regression to the original data); this regression



FIGURE 2.—Relations between length of second dorsal fin and total length. Open circles represent Costa Rican data; solid circles represent Hawaiian data. Solid straight line is linear regression line fitted to Costa Rican data. Broken straight line is linear regression line fitted to Hawaiian data from fish 600 mm. and over in total length. Solid curved line is second degree polynomial fitted to all Hawaiian data.



FIGURE 3.—Relations between length of anal fin and total length. Open circles represent Costa Rican data; solid circles represent Hawaiian data. Solid straight line is linear regression line fitted to Costa Rican data. Broken straight line is linear regression line fitted to Hawaiian data from fish 600 mm. and over in total length. Solid curved line is second degree polynomial fitted to all Hawaiian data.

also is plotted in figure 4. Analysis of covariance shows that the slopes of the two regressions do not differ more than might be expected by chance, but the levels do; the longest dorsal spines of Hawaiian fish appear on the average to be a small, constant percentage shorter than the longest dorsal spines of Costa Rican fish.

Similarly, the logarithms of length of longest dorsal finlet against logarithm of total length yielded a linear regression for the Hawaiian measurements on all sizes of fish, and proved also to provide a good fit to the Costa Rican data for



FIGURE 4.—Relations between length of longest dorsal spine and total length. Open circles and fine line represent Costa Rican data; solid circles and heavy line represent Hawaiian data.

which Schaefer (1948) had fitted a linear regression to the original data. Again, the resulting regressions, plotted in figure 5, when subjected to covariance analysis, indicate a small, constant average percentage difference between finlet lengths of the two populations, the Hawaiian fish having the longer finlets.

Head length and distances from snout to fin insertions

As mentioned earlier, Godsil (1948) has published the measurements of total length, head length, and distances from tip of shout to the insertions of first dorsal, second dorsal, anal, and ventral fins for nearly 2,000 specimens of yellowfin tuna from the American west coast between Cape San Lucas and Panama. The original measurements were published with his analyses of them, so we are able to compare these extensive data both with the Costa Rican data published by Schaefer (1948) and with the Hawaiian data presented herein. In figures 6 to 10 have been plotted head length and distances from snout to fin insertions against total length, which is taken in each case as the independent variable. For each of the three groups of data (Godsil's, Costa

Rican, Hawaiian) have been plotted the mean values of the two variables in each graph for each 10 cm. of total length. To the pooled west-coast data (Godsil's plus my Costa Rican) have been fitted and plotted linear regressions. Also plotted are the curvilinear regressions computed by Godsil (1948, p. 13) for his data, of the type y=a+bx+c/x. On the same graphs have been plotted also the linear-regression line best fitting the Hawaiian data and the best-fitting curvilinear regression of the type selected by Godsil.

For the Hawaiian data, except in one case (snout to insertion of second dorsal of Hawaiian fish), the curvilinear regressions provide a slight improvement in fit over the linear regressions. Inspection of the figures, however, reveals that the differences between the linear and curvilinear regressions are small in comparison with the differences between west-coast and Hawaiian samples. The reduction of the variance about the regression line also is very small in comparison with the difference between the two regions when a curvilinear rather than a linear equation is employed. In consequence, the linear-regression equations will be employed below in considering the application of analysis of covariance to the comparison of samples.



FIGURE 5.—Relations between length of longest dorsal finlet and total length. Open circles and fine line represent Costa Rican data; solid circles and heavy line represent Hawaiian data.



FIGURE 6.—Relations between head length and total length. Solid circles represent Hawaiian data; open circles represent Costa Rican data; solid triangles represent Godsil's west-coast data. Fine solid line is linear regression line fitting west-coast data, while heavy solid line is linear regression line fitting Hawaiian data. Fine broken line is Godsil's curvilinear regression for west-coast data, while heavy broken line is similar regression fitted to Hawaiian data.

A detailed analysis of covariance is not necessary to arrive at the conclusion that with respect to these dimensions the samples from the Hawaiian Islands are different from the samples from the west coast. It is quite obvious from the plots of the mean values for each 10-cm. size class (figs. 6 to 10) that the head length and the distances from snout to the fin insertions are significantly shorter for Hawaiian than for westcoast yellowfin tuna at the larger sizes. If a statement of probability is desired to test a null hypothesis respecting difference between regions, one may proceed in a manner similar to that suggested above in the case of pectoral-fin lengths, confining attention for sake of simplicity to the larger sizes of tuna, say over 800 mm. in total length.

Considering fish of size classes between 800 mm. and 1,600 mm. in total length, for which specimens were available both from the west coast and from Hawaii, the points for the mean values of each 10-cm. length class of Hawaiian fish fall below the values expected on the basis of west-coast data in all cases for head length (fig. 6), snout to insertion of anal (fig. 7), snout to insertion of second dorsal (fig. 8), and snout to insertion of ventral (fig. 9). Since there are 8 such points for each dimension, and under a null hypothesis they might equally well be above or below the value expected from west-coast data, the probability of the observations on the hypothesis is $(\frac{1}{2})^8 = \frac{1}{256}$ for each dimension, which is unlikely. For snout to insertion of first dorsal, one point (900-mm. size class) falls barely above the expected value; the probability of having at most one point above the expected value under the null hypothesis is $(\frac{1}{2})^8 + 8(\frac{1}{2})^8 = \frac{9}{256}$.

By the conventional methods of analysis of covariance (Kendall 1946, p. 237 *et seq.*), we may also test for each of the dimensions the null hypotheses (1) that the sample from the west coast and the sample from Hawaii may both be represented by a single linear-regression equation



FIGURE 7.—Relations between distance from snout to insertion of anal fin and total length. Solid circles represent Hawaiian data; open circles represent Costa Rican data; solid triangles represent Godsil's west-coast data. Fine solid line is linear regression line fitting west-coast data, while heavy solid line is linear regression line fitting Hawaiian data. Fine broken line is Godsil's curvilinear regression for west-coast data, while heavy broken line is similar regression fitted to Hawaiian data.





FIGURE 8.—Relations between distance from snout to insertion of second dorsal fin and total length. Solid circles represent Hawaiian data; open circles represent Costa Rican data; solid triangles represent Godsil's west-coast data. Fine solid line is linear regression line fitting west-coast data, while heavy solid line is linear regression line fitting Hawaiian data. Fine broken line is Godsil's curvilinear regression for west-coast data, while heavy broken line is similar regression fitted to Hawaiian data.



FIGURE 9.—Relations between distance from snout to insertion of ventral fin and total length. Solid circles represent Hawaiian data; open circles represent Costa Rican data; solid triangles represent Godsil's west-coast data. Fine solid line is linear regression line fitting west-coast data, while heavy solid line is linear regression line fitting Hawaiian data. Fine broken line is Godsil's curvilinear regression for west-coast data, while heavy broken line is similar regression fitted to Hawaiian data.

and, if this be false, (2) that the regression coefficients (slopes) of the regression lines fitting the samples from the two regions are equal. As may be seen from the variance ratios computed in table 3, both these hypotheses are to be rejected for each dimension considered, the west-coast data in this table including the measurements of both Schaefer and Godsil. If we compare the Hawaiian data with the data of Schaefer alone (table 4) we find here also that for no character considered may the data from the two regions be represented by a single linear-regression equation. In two cases, however, indicated by footnotes in the table, the appropriate variance ratio indicates that there is not sufficient reason from these particular data to reject the hypothesis of equality of regression coefficients. In general, it is quite apparent that for each character the regression lines are different

for the two regions and that they differ in slope.

Comparison of the regression lines of the dimensions of tuna from different regions is perfectly straightforward so long as we are able to assume that the sample regression lines are representative of the tuna populations of the regions in each case. As has been noted earlier, however, Godsil found that repeated samples from the west coast yielded regression lines (curvilinear) for which a null hypothesis could not be supported. The same thing is true if linear regressions are applied to his data (table 5). His various subgroups along the west coast differ significantly among themselves, and for each dimension they differ in respect of the regression coefficients. As may be seen from table 6, comparison of my Costa Rican data with Godsil's data from Costa Rica alone (his samples 4, 5, and 12) reveals that a single linear-regression equation does not, for any dimension, accurately describe both. It is quite evident that differences may be expected among different samples from the same region. The problem, then, is to determine whether the differences between regions are greater than might reasonably be expected among different samples from the same region. In comparing Hawaiian and west-coast data, where the differences are so large that the distributions of means of subclasses (size groups) are completely separate between the two regions for the most part, the answer is fairly obvious from the graphs of the type herein presented. In table 7 have been tabulated the linear-regression coefficients for each of Godsil's 13 samples, for my Costa Rican sample, and for the Hawaiian sample. From this tabulation it may readily be seen that the Hawaiian regression coefficients fall, for each dimension, well below the lowest value encountered among the several westcoast subsamples.

Although in the case at hand we are spared the need for an efficient means of comparing variation between samples within a region with differences between regions where a null hypothesis is not valid for samples within the region, this will not in general be true. The desirability of a test for application in other, less-clear situations is sufficiently great that some examination of the problem seems warranted, particularly in view of the fact that Godsil (1948) has already attempted to develop and employ such a test. We wish, therefore, to consider the problem of measuring the differences between groups where a null hypothesis is not satisfied.



FIGURE 10.—Relations between distance from snout to insertion of first dorsal fin and total length. Solid circles represent Hawaiian data; open circles represent Costa Rican data; solid triangles represent Godsil's west-coast data. Fine solid line is linear regression line fitting west-coast data, while heavy solid line is linear regression line fitting Hawaiian data. Fine broken line is Godsil's curvilinear regression for west-coast data, while heavy broken line is similar regression fitted to Hawaiian data.

Source of variation	Degrees of freedom	Sum of squares	Mean square	Varlance ratios
Head length: Deviations from total regression Deviations from regressions within regions	2, 158 2, 156	73, 920 32, 698	15. 17	$\frac{20, 611}{15, 17}$ =1359.0
Differences between regions Differences between regression coefficients Differences between adjusted means	2 1 1	41, 222 6, 257 34, 965	20, 611 6, 257	$\frac{6,257}{15,17}$ =412.5
Shout to insertion first dorsal: Deviations from total regression Deviations from regressions within regions	2, 156 2, 154	67, 803 51, 427	23. 88	$\frac{8,188}{23.88}$ =342.9
Differences between regions Differences between regression coefficients Differences between adjusted means	· 2 1 1	. 16, 376 3, 431 12, 945	8, 188 3, 431	$\frac{3,431}{23.88} = 143.7$
Snout to insertion ventral: Deviations from total regression Deviations from regressions within regions	2, 110 2, 108	87, 708 51, 946	24.64	$\frac{17,881}{24.64} = 725.7$
Differences between regions Differences between regression coefficients Differences between adjusted means	2 [.] 1	35, 762 7, 709 28, 053	17, 881 7, 709	$\frac{7,709}{24.64}$ =312.9
Snout to insertion second dorsal: Deviations from total regression Deviations from regressions within regions	2, 156 2, 154	102, 228 77, 765	36. 10	$\frac{12,232}{36,10} = 338.8$
Differences between regions Differences between regression coefficients Differences between adjusted means	2 1 1	24, 463 4, 263 20, 200	12, 232 4, 263	$\frac{4,263}{36,10}$ =118.1
Snout to insertion anal: Deviations from total regression. Deviations from regressions within regions	2, 153 2, 151	128, 518 75, 808	35. 24	$\frac{26,355}{35,24} = 747.9$
Differences between regions Differences between regression coefficients Differences between adjusted means	2 1 1	52, 710 12, 017 40, 693	26, 355 12, 017	$\frac{12,017}{35.24}$ =341.0

TABLE 3.—Comparison of Hawaiian data and pooled American west-coast data by covariance analysis, linear regressions

TABLE 4.—Comparisons of Hawaiian data and Schaefer's Costa Rican data by covariance analysis, linear regressions

Source of variation	Degrees of freedom	Sum of squares	Mean square	Variance ratios
Head length: Deviations from total regression Deviations from regressions within regions	241 239	10, 649 9, 399	39. 33	$\frac{625}{39.33}$ =15.89
Differences between regions. Differences between regression coefficients. Differences between adjusted means.	2 1 1	1, 250 143 1, 107	625 143 1, 107	$\frac{143}{39.33} = 13.64$ $\frac{1,107}{39.76} = 27.84$
Snout to insertion first dorsal: Deviations from total regression. Deviations from regressions within regions	245 243	16, 558 14, 550	5 <mark>9. 88</mark>	$\frac{1,004}{59.88} = 16.77$
Differences between regions Differences between regression coefficients Differences between adjusted means	2 1 1	2, 008 552 1, 456	1, 004 552	$\frac{552}{59.88} = 9.22$
Snout to insertion second dorsal: Deviations from total regression	246 244	35, 349 30, 097	123.35	$\frac{2,626}{123.35}$ =21.29
Differences between regions	2 1 1	5, 252 164 5, 088	2, 626 164 5, 088	$\frac{\frac{104}{123.35}}{\frac{5,088}{123.51}} = 41,20$
Snout to insertion anal: Deviations from total regression Deviations from regression within regions	246 244	36, 130 23, 210	95. 12	$\frac{6,460}{95,12}$ =67.91
Differences between regions Differences between regression coefficients Differences between adjusted means	- <mark>2</mark> - 1 1	12, 920 536 12, 384	6, 460 536	$\frac{536}{95.12}$ =5.63

¹ Not significant.

Source of variation	Degrees of freedom	Sum of squares	Mean square	Variance ratios
Head length: Deviations from total regression Deviations from regression within groups	1, 909 1, 885	23, 049 18, 294	9. 705	$\frac{198.1}{9.705}$ =20.41
Differences among groups. Differences among regression coefficients. Differences among adjusted group means.	24 12 12	4, 755 1, 773 2, 982	198. 1 147. 8	$\frac{147.8}{9.705}$ =15.23
Shout to insertion first dorsal: Deviations from todal regression Deviations from regression within groups	1, 909 1, 885	$rac{36.411}{31.623}$	16.78	$\frac{199.5}{16.78}$ =11.80
Differences among groups Differences among regression coefficients Differences among adjusted group means	24 12 12	4, 788 2, 248 2, 540	199. 5 187. 3	$\frac{187.3}{16.78}$ =11.17
Snout to insertion ventral: Deviations from total regression Deviations from regression within groups	1, 907 1, 883	37, 960 35, 269	18.73	$\frac{112, 1}{18, 73}$ = 5, 99
Differences among groups Differences among regression coefficients Differences among adjusted group means	24 12 12	2, 691 591 2, 100	112. 1 49. 25	$\frac{49,25}{18,73} = 2.63$
Snout to insertion second dorsal: Deviations from total regression Deviations from regression within groups	1,908 1,884	47, 560 40, 871	21.69	$\frac{278.7}{21.69} = 12.85$
Differences among groups Differences among regression coefficients Differences among adjusted group means	24 12 12	6, 689 4, 379 2, 310	278, 7 . 364, 9	$\frac{364.9}{21.69}$ = 16.82
Snout to insertion anal: Deviations from total regression Deviations from regression within groups	1,905 1,881	51, 914 42, 615	22.65	$\frac{387.5}{22.65} = 17.11$
Differences among groups Differences among regression coefficients Differences among adjusted group means	24 12 12	9, 299 2, 745 6, 354	387, 5 228, 8	$\frac{228.8}{22.65} = 10.10$

TABLE 5.—Comparison of subgroups, Godsil's west-coast data, by covariance analysis, linear regressions

TABLE 6.-Comparisons of Schaefer's and Godsil's Costa Rican data by covariance analysis, linear regressions

· .

Source of variation	Degrees of	Sum of	Mean	Variance
	freedom	squares	square	ratios
Head length: Deviations from total regression Deviations from regression within groups		10, 018 9, 730	12.07	$\frac{144}{12.07}$ =11.93
Differences between groups	2	288	144	$\frac{83}{12.07} = 6.88$
Differences between regression coefficients	1	83	83	
Differences between adjusted group means	1	205	205	
Snout to insertion first dorsal: Deviations from total regression Deviations from regression within groups	SÚS 806	16, 550 16, 226	20. 18	$\frac{142}{20.18} = 7.04$
Differences between groups	2	284	142	$\frac{120}{20.18} = 5.95$
Differences between regression coefficients	1	120	120	
Differences between adjusted group means	1	164	164	
Snout to insertion second dorsal: Deviations from total regression Deviations from regression within groups	807 805	27, 293 27, 162	33. 74	$\frac{65.5}{33.74} = 12.00$
Differences between groups	2	131	65, 5	
Differences hetween regression coefficients.	1	60	60	
Differences between adjusted group means	1	71	71	
Snout to insertion anal: Deviations from total regression Deviations from regression within groups	806 804	23,849 23,412	29. 12	$\frac{218}{29, 12} = 7, 49$
Differences between groups.	2	437	218	$\frac{52}{29, 12} = 1.76$ $\frac{385}{29, 15} = 13.22$
Differences between regression coefficients.	1	52	52	
Differences between adjusted group means.	1	385	385	

1 Not significant.

Denote by x_{ij} , y_{ij} the pair of variate values for the *i*th member of the *j*th group, by n_j the number of members of the *j*th group, and by p the number of groups. Also let x_{ij} and y_{ij} be the mean values of the variates in the *j*th group, x_{i} , and y_{i} be the mean values of the variates for the total of all groups, and N be the total of all n_j . The variances about the linear-regression lines may be analyzed as follows:

Variation	Degrees of freedom	Sum of squares	Mean square
Total, from regression b ₀	N-2	$S = \sum_{i_{j} j} (y_{i_{j}} - y_{})^{2} - b_{0} \sum_{i_{j} j} (x_{i_{j}} - x_{}) (y_{i_{j}} - y_{})$	S/N-2
Within groups, from regression b _f	N-2p	$S_{2} = \sum_{i, j} (y_{ij} - y_{.j})^{2} - \sum_{i, j} b_{i} (x_{ij} - x_{.j}) (y_{ij} - y_{.j})$	S ₂ /N-2p=s ₂
Differences between groups	2p-2	$S_{1} = \sum_{i} n_{i}(y_{,i} - y_{,.})^{2} - b_{0} \sum_{i} n_{i}(x_{,i} - x_{,.})(y_{,i} - y_{,.})$ $- \sum_{i,j} (b_{0} - b_{i})(x_{ij} - x_{,j})(y_{ij} - y_{,i})$	S ₁ /2 <i>p</i> -2= <i>s</i> ₁

 TABLE 7.—Regression coefficients for regressions of various dimensions on total length, for samples from the American west coast and Hawaii

· ·	Head length	Snout to insertion first dorsal	Snout to insertion ventral	Snout to insertion second dorsal	Snout to in- sertion anal
Godsil's west-coast samples: No. 1 No. 2 No. 3 No. 5 1 No. 6 No. 7 No. 8 No. 10 No. 11 No. 12 1 No. 13 All samples Schaefer's Costa Rican Tamples	0. 24315 . 27902 . 24339 . 23771 . 24118 . 26280 . 26580 . 26001 . 26014 . 25614 . 25699 . 24356 . 24356 . 24356	0. 27134 . 29256 . 26627 . 25647 . 25743 . 25746 . 28340 . 28341 . 28015 . 26397 . 30858 . 28519 . 26148 . 26346	0. 26520 29840 27189 27185 27210 29487 27615 27536 28676 29405 27348 30068 20858 27348	0. 50285 . 49022 . 43265 . 47464 . 48137 . 50244 . 47767 . 50448 . 50524 . 50524 . 50524 . 50533 . 49191 . 50391 . 50207 . 48358 . 47675	0. 54569 . 55697 . 53736 . 53856 . 54344 . 57659 . 54344 . 55711 . 54410 . 58550 . 55914 . 58009 . 54383 . 53508
Hawanan samples	. 22567	. 24821	. 25259	.46914	. 51941

¹ Samples from Costa Rican waters.

Where b_o is the regression coefficient for all data pooled and b_j is the regression coefficient for the j^{th} group.

When the null hypothesis is satisfied s_1 and s_2 are both unbiased estimates of the variance about the regression line, and their ratio will be distributed in the F distribution.

In the case where the null hypothesis is not satisfied, but a single regression coefficient adequately describes the effect of x on y for all groups, we may subtract

$$Y_{ij} = y_{..} + b_o(x_{ij} - x_{..})$$

from each value of y_{ij} to allow for differences in the *x* variate. The new variable $y_{ij}^{i} = y_{ij} - Y_{ij}$ is completely corrected for variations in *x*, so that differences between adjusted means of groups will be independent of the values of x. We may take, then, an estimate of the differences among the adjusted group means as a measure of the differences between groups which will not be affected by differences in size composition (values of x) of the samples from the different groups (Kendall 1946, p. 244). Geometrically, in this case, the lines are parallel, so that the distance between lines is constant for all values of x.

In the case where a single regression coefficient does *not* represent the effect of x on y for all groups, geometrically where the lines are not parallel, any measurement of the distance between lines will depend on the value or values of x employed for the measurement of the distance. Differences between corrected group means will, then, not be independent of the x values. Geometrically, the distances between regression lines will be dependent upon the selection of the place where the distances are measured. In this situation, obviously, differences between adjusted group means are of small value in measuring differences between groups, when the values of x are selected arbitrarily.

Godsil's statistic (Godsil 1948, p. 9, table 4), the mean-square deviation of the sample regression line of the group from the sample regression line of all data pooled, based on curvilinear regressions, is similarly dependent on the distribution of the xvalues of the variates composing the groups, since the regression coefficients are not equal (the lines are not parallel). Its employment as a standard for judging differences between regions as compared with differences among groups within the region is, therefore, subject to strong objection. It seems, then, that where the groups within a region differ in their regression coefficients, as is true in the present instance, we have no method of measuring with any precision the differences among these groups as a basis of judging whether a further sample from another region could reasonably be expected to belong to the same population as that from which the groups in question were drawn. Of course, in the event the regression coefficient itself is not size-connected, it may be used to characterize the group, and one might compare the variation among group regression coefficients with the observed value of the regression coefficient from the further sample from another region (e. g. table 7).

Pending development of a method of precise analysis, comparison of differences among regression lines within regions with differences between regions does not appear to be very fruitful, except in those cases where the difference between regions is so very much greater than differences among samples within a region that it is quite apparent from a simple graph of the data and no precise method of analysis is required.

As a practical procedure it appears best, perhaps, to select fish from each region from many different schools, and of sizes that will cover the entire range available, and then, in comparing data between regions by covariance analyses, to compare samples of similar size range. In this manner any variation between groups within the region will tend to be assimilated into the variance of the total sample for the whole region, and the total sample will be nearly representative of the population of the region.

Other dimensions

Comparison of the regression of diameter of iris on head length of Hawaiian specimens with that of Costa Rican specimens indicates that the relation is different in the two regions. The relations and the means of the two variates for each 10 centimeters of total length are plotted in figure 11.

Comparison of Hawaiian and Costa Rican data respecting regressions of length of maxillary on head length, body depth on total length, and weight on total length indicated that in each case the two samples might have been drawn at random from a single population so far as these characters are concerned.

Counts of gill rakers

Counts of total gill rakers of 188 Hawaiian tuna (table 1) have a mean value of 29.66 with a standard error of .0870. Schaefer's (1948) Costa Rican data on 45 specimens have a mean value of 30.60 with a standard error of .186, while Godsil and Byer's (1944) counts of 60 American-west-coast specimens have a mean of 30.35 with a standard error of .146. Comparison of the Costa Rican and Godsil and Byer's data yields a t value of 1.06, so that the null hypothesis is reasonable and we may pool these data to estimate the mean gillraker count of yellowfin from the American west coast as 30.46 with a standard error of .116. The difference of .80 between this value and the Hawaiian mean is associated with a t value of 5.52.

We have verified from our Hawaiian data that there is no correlation between size of fish and gillraker count. This character seems to offer good possibilities for racial analysis of tunas for that reason, since it will avoid the difficulties in comparisons which plagued us in regression analyses.

DISCUSSION

Hawaiian yellowfin tuna differ from those of the American west coast in having, on the average, longer pectoral fins at the same fish size, and this difference is greater for the larger fish. The same is true of the second dorsal and anal fins, but in these cases the fins of the Hawaiian fish also grow at an accelerated rate compared to west-coast fish, so that the difference in fin lengths among the largest fish sizes is very striking. The first dorsal spine appears to be consistently shorter among Hawaiian fish, while the longest dorsal finlet is longer.

Among Hawaiian fish, the distance from tip of snout to the posterior edge of the opercle and to the various fin insertions increases, relative to total length, more slowly than among west-coast fish so that all these dimensions are shorter, on the average, for the large fish from Hawaii than for west-coast fish of comparable size. From this it is evident that the posterior part of the trunk grows faster among Hawaiian fish so that at large sizes, say above 700 or 800 mm., the posterior part of the body is more elongate than among westcoast fish of similar sizes.

On the basis of the magnitude and consistency of these differences between the biometric charac-



FIGURE 11.—Relations between diameter of iris and length of head. Open circles and fine line represent Costa Rican data; solid circles and heavy line represent Hawaiian data.

teristics of yellowfin tuna from the Hawaiian Islands and from the American west coast, there is no doubt that these two populations are to be regarded as distinct. The possibility of some mixing between them is not excluded, but if any exists it must be sufficiently small to permit the two populations to maintain their characteristic differences.

The statistical comparison of body-proportion data on tunas from different regions by regression analysis is beset with difficulties which are beyond the scope of this paper to deal with, and which seem not to be critical in this instance where the differences dealt with are of sufficient magnitude that sensitive methods are not required. The problem merits, however, further attention since it will become acute where differences to be measured are small. This problem may be avoided by employing denumerable characters which are not size-connected. Gill-raker counts seem to be a useful character of this sort. The Hawaiian and westcoast yellowfin-tuna populations are quite distinct with respect to mean gill-raker count.

The fact, brought out by this study, that the yellowfin tuna of the central Pacific belong to a population distinct from that along the American west coast, has important implications in the development and management of the tuna fisheries. Since the yellowfin tuna of these regions belong to different populations which do not freely intermix, a fishery on one can have no effect on the abundance of the other. The fishery along the west coast is not tapping the entire yellowfintuna resource of the Pacific.

The various biometric differences demonstrated herein are of about the same magnitude as the differences between yellowfin tuna from the waters of the American west coast and from the Atlantic off Africa (Schaefer and Walford 1950). In some cases, such as the lengths of second dorsal and anal fins, the differences between the two samples from the Pacific are even more striking than the differences between African and American west-coast samples. If it is borne out by further study that the variation within oceans is about as great as the variation between them, it will be necessary to regard all the yellowfin tunas as belonging to a single species. It is particularly desirable that a series of specimens be examined from the Indian Ocean, whence comes the type of N. argentivittatus, which has priority among the several descriptions of species of Neothunnus, in order to settle the question of nomenclature.

LITERATURE CITED

FRASER-BRUNNER, A.

1950. The fishes of the family Scombridae. Ann. Mag. Nat. Hist., Ser. 12, vol. 3, pp. 131-163, 35 figs.

GODSIL, H. C.

1948. A preliminary population study of the yellowfin tuna and the albacore. California Div. Fish and Game, Fish. Bull. No. 70, 90 pp., 7 tables, 22 figs. GODSIL, H. C., and R. D. BYERS

1944. A systematic study of the Pacific tunas. California Div. Fish and Game, Fish. Bull. No. 60, 131 pp., 18 tables, 76 figs.

JUNE, FRED C.

1950. Preliminary fisheries survey of the Hawaiian-Line Islands area. Part I—The Hawaiian long-line fishery.
U. S. Fish and Wildlife Service Com. Fish. Rev., vol.
12, No. 1, pp. 1–23, 8 tables, 18 figs.

KENDALL, MAURICE G.

1946. The advanced theory of statistics. Vol. II. 521 pp. Charles Griffin & Co., London.

MARR, JOHN C., and MILNER B. SCHAEFER

1949. Definitions of body dimensions used in describing tunas. U. S. Fish and Wildlife Service, Fishery Bull. No. 47, vol. 51, pp. 241-244, 1 fig.

MARTIN, W. R.

1949. The mechanics of environmental control of body form in fishes. Univ. Toronto Studies, Biol., No. 58, Pub. Ont. Fish. Res. Lab., No. 70, 72 pages, 13 tables, 24 figs.

SCHAEFER, MILNER B.

1948. Morphometric characteristics and relative growth of yellowfin tunas (*Neothunnus macropterus*) from Central America. Pacific Science, vol. 2, No. 2, pp. 114-120, 2 tables.

SCHAEFER, MILNER B., and LIONEL A. WALFORD

1950. Biometric comparison between yellowfin tunas (*Neothunnus*) of Angola and of the Pacific Coast of Central America. U. S. Fish and Wildlife Service, Fish. Bull. No. 56, vol. 51, pp. 425-443, 3 tables, 15 figs.

0